

Exercise 1 $X_n = \sum_{i=1}^n Y_i$

$$\mathbb{P}(X_n \geq a) \leq \frac{\mathbb{E}[e^{\mu X_n}]}{\mu^a} = \frac{\mathbb{E}\left[\prod_{i=1}^n e^{\mu Y_i}\right]}{\mu^a} = \frac{\prod_{i=1}^n \mathbb{E}[e^{\mu Y_i}]}{\mu^a} = \frac{\mathbb{E}[e^{\mu Y_1}]^n}{\mu^a} \quad \forall \mu > 1$$

Chernoff bounds
(Y_i)_i iid.

The other inequality is similar.

Exercise 2 $Y_i = \begin{cases} 1 & \text{w.p. } 1/2 \\ -1 & \text{w.p. } 1/2 \end{cases} \quad X_n = \sum_{i=1}^n Y_i$

$$\mathbb{P}(X_n \geq a) = \mathbb{P}(X_n \leq -a) \leq \frac{\mathbb{E}[e^{\mu Y_1}]^n}{\mu^a} = \frac{\mathbb{E}[e^{\lambda Y_1}]^n}{e^{a\lambda}} = \frac{\left(\frac{e^\lambda + e^{-\lambda}}{2}\right)^n}{e^{a\lambda}}$$

$\mu = e^\lambda$ with $\lambda > 0$

$$\leq \frac{(e^{\lambda/2})^n}{e^{a\lambda}} = e^{\frac{\lambda}{2}n - a\lambda} = e^{\frac{n}{2}\frac{\lambda^2}{n^2} - \frac{a^2}{n}} = e^{-\frac{a^2}{2n}} \quad (1)$$

$$\frac{e^\lambda + e^{-\lambda}}{2} \leq e^{\lambda/2}$$

$\hookrightarrow \lambda = \frac{a}{n}$ (min of $\frac{\lambda^2}{2} - a\lambda$)

since $e^x = 1 + x + \frac{x^2}{2} + \frac{x^3}{3!} + \dots$

$$\Rightarrow \frac{e^\lambda + e^{-\lambda}}{2} = 1 + \frac{\lambda^2}{2} + \frac{\lambda^4}{4!} + \frac{\lambda^6}{6!} + \dots + \frac{\lambda^{2k}}{(2k)!}$$

$$e^{\lambda/2} = 1 + \frac{\lambda}{2} + \frac{\lambda^2}{2! \cdot 2^2} + \frac{\lambda^3}{3! \cdot 2^3} + \dots + \frac{\lambda^{2k}}{k! \cdot 2^k}$$

we conclude using that $2^k k! \leq (2k)!$

Corollary: $Z_i = \begin{cases} 0 & p=1/2 \\ 1 & p=1/2 \end{cases} \quad S_n = \sum_{i=1}^n Z_i \quad \mathbb{E}[S_n] = n/2$

$$\mathbb{P}\left(S_n - \frac{n}{2} > a\right) = \mathbb{P}\left(\sum_{i=1}^n (Z_i - 1/2) > a\right) = \mathbb{P}\left(\sum_{i=1}^n \underbrace{(2Z_i - 1)}_{\substack{-1 \text{ w.p. } 1/2 \\ 1 \text{ w.p. } 1/2}} > 2a\right) <$$

$$< e^{-\frac{(2a)^2}{2n}} = e^{-\frac{2a^2}{n}}$$

(1)

Exercise 3

1. Observe that by taking σ and reverse of σ , we can have
of upsets $\leq \frac{1}{2} \binom{n}{2}$

2. Fix $\sigma: [n] \rightarrow [n]$, T_n is a uniform tournament, $U_n = \#$ of upset in (T_n, σ) .
 $U_n = \sum_{i+j} Y_{ij}^n$, where $Y_{ij}^n = \begin{cases} 1 & \text{if } (i, j) \text{ is an upset in } T_n \\ 0 & \text{otherwise} \end{cases}$

It's easy to show that $\mathbb{P}(Y_{ij}^n = 0) = \mathbb{P}(Y_{ij}^n = 1) = \frac{1}{2}$. Note that $\mathbb{E}[U_n] = \frac{1}{2} \binom{n}{2}$
and so

$$\mathbb{P}\left(U_n - \frac{1}{2} \binom{n}{2} < -a\right) \leq e^{-2a^2 / \binom{n}{2}} = e^{-\frac{2a^2}{n(n-1)/2}} < e^{-\frac{4a^2}{n^2}}$$

3. We set $a = n^{3/2} \sqrt{\log(n)}$ (I want $e^{-\frac{4a^2}{n^2}} < n!$
 $\Rightarrow a = n^{3/2} \sqrt{\log n}$) obtaining

$$\mathbb{P}\left(U_n < \frac{1}{2} \binom{n}{2} - n^{3/2} \sqrt{\log(n)}\right) < e^{-\frac{4}{n^2} n^3 \log(n)} = e^{-4n \log(n)} = n^{-4n}$$

and so

$$\mathbb{P}\left(\exists \sigma \text{ s.t. } U_n(\sigma) < \frac{1}{2} \binom{n}{2} - n^{3/2} \sqrt{\log(n)}\right) \leq \underbrace{n!}_{\substack{\# \text{ of} \\ \text{perm. of size } n}} n^{-4n}$$

4. Using the Stirling's approximation $n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$ we have

$$n! n^{-4n} \sim \sqrt{2\pi n} \frac{n^n}{e^n} n^{-4n} = \sqrt{2\pi n} \frac{n^{-3n}}{e^n} \rightarrow 0.$$

$\Rightarrow n! n^{-4n} \ll 1 \Rightarrow \exists$ a tournament T on n vertices s.t. for every possible ordering σ , the number of upsets in T is at least $\frac{1}{2} \binom{n}{2} - n^{3/2} \log(n)$.