

Exercise 1

$$2. X_n = \sum_{i=1}^{4n} \mathbb{1}_{\{\sigma(i) \equiv i \pmod{n}\}} \quad \text{where } \sigma \text{ is a unif. perm. of size } 4n.$$

$$3. \mathbb{E}[(X_n)_r] = \sum_{\substack{i_1, \dots, i_r \in [4n] \\ \text{distinct}}} \mathbb{E}[\mathbb{1}_{\{\sigma(i_1) \equiv i_1 \pmod{n}\}} \cdots \mathbb{1}_{\{\sigma(i_r) \equiv i_r \pmod{n}\}}] \\ = \mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_r) \equiv i_r \pmod{n})$$

$$4. \mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_r) \equiv i_r \pmod{n}) = \\ = \mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}) \cdot \mathbb{P}(\sigma(i_2) \equiv i_2 \pmod{n} \mid \sigma(i_1) \equiv i_1 \pmod{n}) \cdots \\ \otimes \cdots \mathbb{P}(\sigma(i_r) \equiv i_r \pmod{n} \mid \sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_{r-1}) \equiv i_{r-1} \pmod{n})$$

Now

$$\mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}) = \frac{4}{4n}$$

$$\mathbb{P}(\sigma(i_2) \equiv i_2 \pmod{n} \mid \sigma(i_1) \equiv i_1 \pmod{n}) = \begin{cases} \frac{4}{4n-1} & \text{if } i_1 \not\equiv i_2 \pmod{n} \\ \frac{3}{4n-1} & \text{if } i_1 \equiv i_2 \pmod{n} \end{cases}$$

$$\Rightarrow \mathbb{P}(\sigma(i_2) \equiv i_2 \pmod{n} \mid \sigma(i_1) \equiv i_1 \pmod{n}) \leq \frac{4}{4n-1}$$

Similarly, the factor

$$\mathbb{P}(\sigma(i_j) \equiv i_j \pmod{n} \mid \sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_{j-1}) \equiv i_{j-1} \pmod{n}) \leq \frac{4}{4n-j+1} \quad \left( \begin{array}{l} \text{with equality} \\ \text{when the } i_j \text{'s are} \\ \text{not congruent} \end{array} \right)$$

and so

$$\mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_r) \equiv i_r \pmod{n}) \leq \frac{4^r (4n-r)!}{(4n)!}$$

$$5. \text{ Note that the \# of tuples } (i_1, \dots, i_r) \in [4n]^r \text{ s.t. some } i_j \text{'s are congruent} \\ \text{mod } n \text{ is } \leq \underbrace{\binom{4n}{r-1}}_{\text{choices}} \cdot \underbrace{3(r-1)}_{\text{choices for}} = O(n^{r-1}).$$

$$\text{mod } n \text{ is } \leq \underbrace{\binom{4n}{r-1}}_{\substack{\text{choices} \\ \text{for the first} \\ r-1 \text{ elements}}} \cdot \underbrace{3^{r-1}}_{\substack{\text{choices for} \\ \text{the last element} \\ \text{(congruent to one} \\ \text{of the previous)}}} = O(n^{r-1}).$$

$$\begin{aligned}
 6. \quad \mathbb{E}[(X_n)_r] &= \sum_{\substack{i_1, \dots, i_r \in [4n] \\ \text{distinct}}} \mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_r) \equiv i_r \pmod{n}) \\
 &= \sum_{\substack{i_1, \dots, i_r \in [4n] \\ \text{distinct} \\ \text{and with no} \\ \text{congruent } i_j\text{'s}}} \underbrace{\mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_r) \equiv i_r \pmod{n})}_{= \frac{4^r (4n-r)!}{(4n)!}} \\
 &\quad 4n(4n-4) \dots (4n-4(r-1)) \\
 &+ \sum_{\substack{i_1, \dots, i_r \in [4n] \\ \text{distinct} \\ \text{and with} \\ \text{congruent } i_j\text{'s}}} \underbrace{\mathbb{P}(\sigma(i_1) \equiv i_1 \pmod{n}, \dots, \sigma(i_r) \equiv i_r \pmod{n})}_{\leq \frac{4^r (4n-r)!}{(4n)!}} \\
 &\quad O(n^{r-1})
 \end{aligned}$$

$$\Rightarrow \mathbb{E}[(X_n)_r] \sim 4n(4n-4)(4n-4 \cdot 2) \dots (4n-4(r-1)) \frac{4^r (4n-r)!}{(4n)!}$$

$$\Rightarrow \mathbb{E}[(X_n)_r] \xrightarrow{n \rightarrow \infty} 4^r$$

↑ r-th factorial moment of Poisson(4)

Applying the moment method we can conclude that  $X_n \xrightarrow{d} \text{Poisson}(4)$ .

## Exercise 2

We first check that  $f_X(x)$  and  $f_Y(x)$  are well-def prob. densities.

$$\int_0^{\infty} (1 + \sin(2\pi \log(x))) \frac{1}{\sqrt{2\pi} x} e^{-\log(x)^2/2} dx = \int_{-\infty}^{\infty} (1 + \sin(y)) \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy$$

$\begin{cases} \log(x) = y \\ x = e^y \\ dx = dy e^y \end{cases}$

$$\begin{aligned}
& \int_0^{\infty} (1 + \sin(2\pi \log(x))) \frac{1}{\sqrt{2\pi} x} e^{-\log(x)^2/2} dx \\
&= \int_{-\infty}^{+\infty} (1 + \sin(2\pi y)) \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy \\
&= \underbrace{\int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy}_{= 1 \text{ (Gaussian)}} + \underbrace{\int_{-\infty}^{+\infty} \sin(2\pi y) e^{-y^2/2} dy}_{= 0 \text{ (since } \sin(-x) = -\sin(x))}
\end{aligned}$$

dx = dy e^y

We now show that  $\mathbb{E}[X^p] - \mathbb{E}[Y^p] = 0$

$$\begin{aligned}
\mathbb{E}[Y^p] - \mathbb{E}[X^p] &= \int_0^{\infty} x^p \frac{e^{-\log(x)^2/2}}{\sqrt{2\pi} x} \sin(2\pi \log(x)) dx \\
\log(x) = y \quad \downarrow &= \int_{-\infty}^{+\infty} e^{yp} \frac{e^{-y^2/2}}{\sqrt{2\pi}} \sin(2\pi y) dy \\
&= e^{p^2/2} \int_{-\infty}^{+\infty} \sin(2\pi y) e^{-(y-p)^2/2} dy = 0 \\
&\quad \parallel \sin(2\pi(y-p)) \quad (\sin \text{ is periodic})
\end{aligned}$$